

RESEARCH ARTICLE

Using Google Health Trends to investigate COVID-19 incidence in Africa

Alexander Fulk¹✉, Daniel Romero-Alvarez^{1,2,3}✉*, Qays Abu-Saymeh¹, Jarron M. Saint Onge^{4,5}, A. Townsend Peterson^{1,2}, Folashade B. Augusto¹

1 Department of Ecology & Evolutionary Biology, University of Kansas, Lawrence, Kansas, United States of America, **2** Biodiversity Institute, University of Kansas, Lawrence, Kansas, United States of America, **3** One Health Research Group, Facultad de Medicina, Universidad de las Américas, Quito, Ecuador, **4** Department of Population Health, University of Kansas Medical Center, Kansas City, Kansas, United States of America, **5** Department of Sociology, University of Kansas, Lawrence, Kansas, United States of America

✉ These authors contributed equally to this work.

* daromero88@gmail.com



OPEN ACCESS

Citation: Fulk A, Romero-Alvarez D, Abu-Saymeh Q, Saint Onge JM, Peterson AT, Augusto FB (2022) Using Google Health Trends to investigate COVID-19 incidence in Africa. PLoS ONE 17(6): e0269573. <https://doi.org/10.1371/journal.pone.0269573>

Editor: Shawky M Aboelhadid, Beni Suef University Faculty of Veterinary Medicine, EGYPT

Received: March 7, 2022

Accepted: May 23, 2022

Published: June 7, 2022

Copyright: © 2022 Fulk et al. This is an open access article distributed under the terms of the [Creative Commons Attribution License](https://creativecommons.org/licenses/by/4.0/), which permits unrestricted use, distribution, and reproduction in any medium, provided the original author and source are credited.

Data Availability Statement: All relevant data are within the manuscript and its [Supporting information](#) files.

Funding: YES Folashade Augusto Grant number DMS 2028297. National Science Foundation. <https://www.nsf.gov/> The sponsors did not play any role in the study design, data collection and analysis, or decision to publish, or preparation of the manuscript.

Competing interests: The authors have declared that no competing interests exist.

Abstract

The COVID-19 pandemic has caused over 500 million cases and over six million deaths globally. From these numbers, over 12 million cases and over 250 thousand deaths have occurred on the African continent as of May 2022. Prevention and surveillance remains the cornerstone of interventions to halt the further spread of COVID-19. Google Health Trends (GHT), a free Internet tool, may be valuable to help anticipate outbreaks, identify disease hotspots, or understand the patterns of disease surveillance. We collected COVID-19 case and death incidence for 54 African countries and obtained averages for four, five-month study periods in 2020–2021. Average case and death incidences were calculated during these four time periods to measure disease severity. We used GHT to characterize COVID-19 incidence across Africa, collecting numbers of searches from GHT related to COVID-19 using four terms: ‘coronavirus’, ‘coronavirus symptoms’, ‘COVID19’, and ‘pandemic’. The terms were related to weekly COVID-19 case incidences for the entire study period via multiple linear and weighted linear regression analyses. We also assembled 72 variables assessing Internet accessibility, demographics, economics, health, and others, for each country, to summarize potential mechanisms linking GHT searches and COVID-19 incidence. COVID-19 burden in Africa increased steadily during the study period. Important increases for COVID-19 death incidence were observed for Seychelles and Tunisia. Our study demonstrated a weak correlation between GHT and COVID-19 incidence for most African countries. Several variables seemed useful in explaining the pattern of GHT statistics and their relationship to COVID-19 including: log of average weekly cases, log of cumulative total deaths, and log of fixed total number of broadband subscriptions in a country. Apparently, GHT may best be used for surveillance of diseases that are diagnosed more consistently. Overall, GHT-based surveillance showed little applicability in the studied countries. GHT for an ongoing epidemic might be useful in specific situations, such as when countries have significant levels of infection with low variability. Future studies might assess the algorithm in different epidemic contexts.

Introduction

Coronavirus disease 2019 (COVID-19) is a respiratory disease caused by the severe acute respiratory syndrome coronavirus-2 (SARS-CoV-2) discovered in China in 2019. People infected experience a range of symptoms including headache, fever, difficulty breathing, and loss of taste and smell, or may be completely asymptomatic [1]. Since its discovery, SARS-CoV-2 has spread around the globe, with over 500 million confirmed cases as of May 2022, according to John Hopkins University [2,3]. The elderly (>65 years old), as well as those with pre-existing comorbidities, have the highest risk of mortality if infected [4]. COVID-19 spreads via respiratory particles, which allows it to infect others via contaminated aerosols and droplets suspended in the air in closed spaces [5]. Asymptomatic carriers account for a significant amount of secondary transmissions, with some reports showing that ~80% of infections may occur without symptoms, constituting the source of most secondary COVID-19 cases [1,6,7].

After the large-scale Ebola outbreak in 2015, African leaders were aware that swift and decisive action was needed to avert broad spread of COVID-19 and prevent healthcare system collapse. This awareness led to wide adoption of mitigation and control efforts that circumvented an overwhelming first epidemic wave with a partially structured continental response [8]. Regardless, testing in Africa has been limited: about 75% of COVID-19 diagnoses came from tests conducted in only 10 countries [5,8]. The emergence of new SARS-CoV-2 variants (e.g., Beta, Delta, Omicron, etc) has made it difficult to predict wave dynamics across the continent, echoing other regional contexts [9]. Finally, although vaccination campaigns have been promoted as the definitive solution to the pandemic [10], several countries have struggled with vaccine uptake [11]. Africa as a continent has received only about 6% of the roughly 9 billion doses manufactured so far, even though about 17% of the world's population lives there [12]. Further, uptake has been limited, as only 10% of Africans have been fully vaccinated [13,14]. Factors determining this context include disinformation via social media [15], lack of syringes, lack of health workers to administer vaccines, especially in rural areas [13], limited government planning and testing, and complicated efforts of vaccine allocation [16].

Given difficulties in obtaining accurate and timely data on case counts and other epidemiological metrics for COVID-19 worldwide [17,18], the current pandemic represents an opportunity to use digital epidemiology tools to fill gaps in information. Infodemiology is an area of epidemiology that uses digital data to gain insight into disease dynamics [19,20]. The digital data used for this method of surveillance vary widely and may or may not have not been intended for epidemiological purposes, coming from unexpected sources such as restaurant receipts, Facebook posts, or Google search queries [21–23].

Google developed two specific algorithms to address infectious diseases, Google Flu Trends (GFT) in 2009 and Google Dengue Trends (GDT) in 2011 [21], which, after inquiries into their usefulness, were shut down in 2015 [24]. Currently, Google maintains two portals by which to harvest search query data: Google Trends (GT) and Google Health Trends (GHT). GT inquiries yield a ranked score from 0 to 100 based on the highest frequency of searches for a term in a particular time period. GHT provides search counts from a relative proportion of a random sample of the overall Google search dataset for any particular term in a selected time interval [25]. Both of these portals have limitations, such as possibly excluding certain groups (e.g., the elderly, rural residents, low income populations), lack of detail on who is searching certain terms, and no insight into the underlying motivations of the searches [26].

Digital tools have been used in many instances to predict disease incidence [27–29] including COVID-19. Kurian et al. (2020) evaluated the applicability of GT in predicting COVID-19 cases in the United States (U.S.) in a state-by-state analysis [30]. They found that certain

keywords had a strong correlation with COVID-19 cases, and concluded that GT may be a useful tool for predicting COVID-19 outbreaks. Brodeur et al. (2021) used GT to see how lockdowns affected the well-being in the U.S. [31]. Once lockdowns were implemented, well-being likely decreased, as searches for certain terms such as ‘stress,’ ‘suicide,’ and ‘worry’ increased over the lockdown period. Ahmad et al. (2020) used gastrointestinal-related symptom search terms to determine whether GT could predict COVID-19 incidence, and found correlations between the search terms and increases of COVID-19 cases in multiple regions across the U.S. with a four-week lag [32].

Here, we explored whether GHT search query data correlate with COVID-19 incidence at the country level in Africa, as a potential complementary source for more customary forms of COVID-19 surveillance. We decided to use GHT instead of GT given the semi-quantitative nature of the information recovered by GHT. We collected case and death data for 54 African countries, and used four COVID-19-related search terms (see below) for each country. We then assessed whether Internet access, demography, economic information, or health variables, were associated with GHT searches. Lastly, we calculated a standardized volatility index to illuminate whether variability in the signal of case incidence led to less accurate predictions by GHT.

Methods

COVID-19 incidence data

Daily COVID-19 new cases and death counts were obtained for all 54 African countries from 2 February 2020 to 25 September 2021. Country-level case data were obtained via the Johns Hopkins COVID-19 global time series on the pandemic [33]; data were constrained to laboratory-confirmed cases only. We explored the progression of average daily COVID-19 case and death incidence per 100,000 people in Africa in four time periods, each roughly five months (~150 days) long: (a) 2 February 2020 to 30 June, (b) 1 July to 30 November, (c) 1 December to 30 April 2021, and (d) 1 May to 25 September 2021. We then converted daily new cases into weekly new cases for each of the countries to match the weekly GHT data up to 25 September 2021, for a total of 86 observations. We calculated weekly incidence rates by dividing the number of cases per week by the total population per country in millions [34]. Country-level population data were collected from the forecasted midyear 2020 estimates from the U.S. Census Bureau [35].

Ethics

Human data included in this study was collected through publicly available repositories of anonymized COVID-19 case counts from 54 African countries [1,2]. Thus, the present research does not need a revision by a bioethics committee.

Google Health Trends data

We downloaded data corresponding to four English terms from the GHT application programming interface (API): ‘coronavirus,’ ‘coronavirus symptoms,’ ‘COVID19,’ and ‘pandemic’. Although the four terms are related conceptually, they have the potential to capture a broad spectrum of information specifically related to the studied disease, avoiding non-informative data from less specific words, as has been previously demonstrated [25,36]. We addressed potential language barriers by collecting data for the latter two terms in French and Portuguese. The former two search terms were spelled the same in French and Portuguese, aside from accents, so the English versions of those terms captured a majority of individuals

searching those terms in those other languages. We matched the relative search proportions of these words—which is the raw output provided by GHT [25]—with the weekly COVID-19 case incidence for the selected time period.

Statistical analysis

We used a multiple linear regression model fitted with the four GHT English search terms as independent variables of COVID-19 incidence at the country level for each of the 54 African countries being evaluated. We then performed the same analysis, substituting the latter two terms for their equivalents in French or Portuguese, based on the official or spoken language as determined by Nations Online [37]. The primary outcome measure was the adjusted R^2 statistic as a measure of the best-fitting models. We recorded the largest adjusted R^2 value (in absolute value) from the models with all English or English and French/Portuguese terms. We assumed that an R^2 value >0.5 was the minimum threshold to show relevant associations. If one or more of the four terms chosen did not retrieve search counts from GHT, it was removed from the analysis for that country. At least two terms were included for each region. While multicollinearity may exist in our time series, it will minimally affect the goodness-of-fit as we rely on R^2 instead of p -values for interpreting the association between dependent and independent variables [38]. Finally, to address possible autocorrelation and heteroskedasticity issues in our time series, we performed first-order differencing and ran another analyses with a weighted least squares regression model, giving larger weight to those observations with lower variance. We also recorded the results from this weighted regression as a more conservative measure.

Next, we used the adjusted R^2 statistics collected from the 54 African countries as a dependent variable to explore whether different socio-economic variables might explain the patterns. This analysis was conducted separately for the adjusted R^2 statistics collected from the basic fitted regression and the weighted regression models, respectively. Socio-economic variables for the African countries included Internet access, demographic, economic, and health indicators (Table 1); data were gathered from World Bank [39]. We explored logarithmic transformations of each of these variables to determine whether normalization of the indicators led to stronger correlations. We also included a standardized volatility score calculated using the standardized normalized case incidence data of each country as follows:

$$\text{Volatility} = \frac{1}{n-1} \sum_{i=2}^n |Y_i - Y_{i-1}|$$

in which n is the total number of observations and Y is the normalized case incidence per country. The average of the absolute difference (i.e., volatility) summarizes the COVID-19 case incidence signal, reflecting if it is relatively constant or fluctuates broadly from week to week [25]. Overall, we explored a total of 72 potential explanatory variables (Table 1 and S1 Table).

Variables were analyzed individually using a pair-wise univariate linear regression and collectively in a multivariate stepwise regression, in which variables were added and removed iteratively to obtain a subset of variables providing the best model outcome according to the Akaike Information Criterion (AIC). In addition, variables were analyzed using a least absolute shrinkage and selection operator (i.e., LASSO) regression for both untransformed and log-adjusted data to avoid overfitting and produce simpler models. Countries with missing variable information were removed from the univariate regression with that particular variable (38/72; 53% of variables had at least one country removed, S1 Table), and only variables with information for every country were used in the stepwise and LASSO regressions. All analyses were done for both adjusted R^2 values collected from the basic regression and the weighted regression. All analyses were performed in R [40]. Data and scripts to replicate the results of

Table 1. Independent variables explored in the present study. Different categories were selected based on their perceived potential to explain patterns of Google Health Trends and COVID-19 regression models. We also evaluated the log of each variable, for a total of 72 variables.

Category	Indicator	
Internet access	1. Percentage of population with access to electricity.	
	2. Fixed total number of broadband subscriptions in a country.	
	3. Fixed broadband subscriptions per 100 people.	
	4. Fixed total number of telephone subscriptions in a country.	
	5. Fixed telephone subscriptions per 100 people.	
	6. Percentage of individuals using the Internet.	
	7. Total number of mobile cellular subscriptions in a country.	
	8. Mobile cellular subscriptions per 100 people.	
	9. Secure Internet servers per 1 million people.	
Demographics	10. Percentage of people 15 years and above that are literate.	
	11. Percentage of people using at least basic drinking water services.	
	12. Percentage of people using at least basic sanitation services.	
	13. Percentage of people using safely managed drinking water services.	
	14. Percentage of people using safely managed sanitation services.	
	15. Percentage of people with basic hand washing facilities.	
	16. Total population.	
	17. Population density as people per square km of land area.	
	18. Total urban population.	
Economics	19. Percentage of urban population.	
	20. Percentage of GDP* for current health expenditure.	
	21. GDP* (current \$ U.S. value).	
Health	22. GDP* per capita (current \$ U.S. value).	
	23. Average weekly cases over the studied period.	
	24. Community health workers per 1,000 people.	
	25. Cumulative total deaths over the study period.	
	26. Hospital beds per 1,000 people.	
	27. Total life expectancy (years) at birth.	
	28. Nurses and midwives per 1,000 people.	
	29. Physicians per 1,000 people.	
	30. Percentage of population 15–49 years with HIV.	
	31. Prevalence of moderate or severe food insecurity in the population.	
	32. Prevalence of severe food insecurity in the population.	
	33. Percentage of people at risk of catastrophic expenditure for surgical care.	
	34. Percentage of people at risk of impoverishing expenditure for surgical care.	
	35. Smoking prevalence for people above 15 years.	
	Case fluctuation	36. Volatility score for a country calculated using weekly incidence.

*GDP = gross domestic product; HIV = human immunodeficiency virus. Raw values of the variables can be found in S1 Table.

<https://doi.org/10.1371/journal.pone.0269573.t001>

this study are available in a GitHub repository accompanying this publication (<https://github.com/alxjfulk/GHT-and-COVID19-code>).

Results

Examining the distribution of first cases among the 54 African countries, we observed that dates of first reported COVID-19 cases were centered around March 2020. Egypt (EGY)

reported the first case of COVID-19 on the continent on 14 February 2020, 15 days after the World Health Organization (WHO) declared the COVID-19 epidemic an emergency of international concern [41]. Comoros (COM) and Lesotho (LSO) were the last countries to report COVID-19 introductions, with first cases on 30 April and 13 May 2020, respectively (Fig 1).

Countries with highest COVID-19 case incidences for the first time period include Djibouti (3.39 cases per 100,000 people), São Tome and Principe (2.25), and South Africa (1.79) (Fig 2). During the second period, Cameroon (10.7), Libya (7.78), and South Africa (7.39) were most affected (Fig 2). For the third and fourth periods, countries across the continent reported increased COVID-19 incidences, with Seychelles (third period = 39.1; fourth period = 109), Tunisia (third period = 12.0; fourth period = 22.8), Botswana (third period = 10.3; fourth

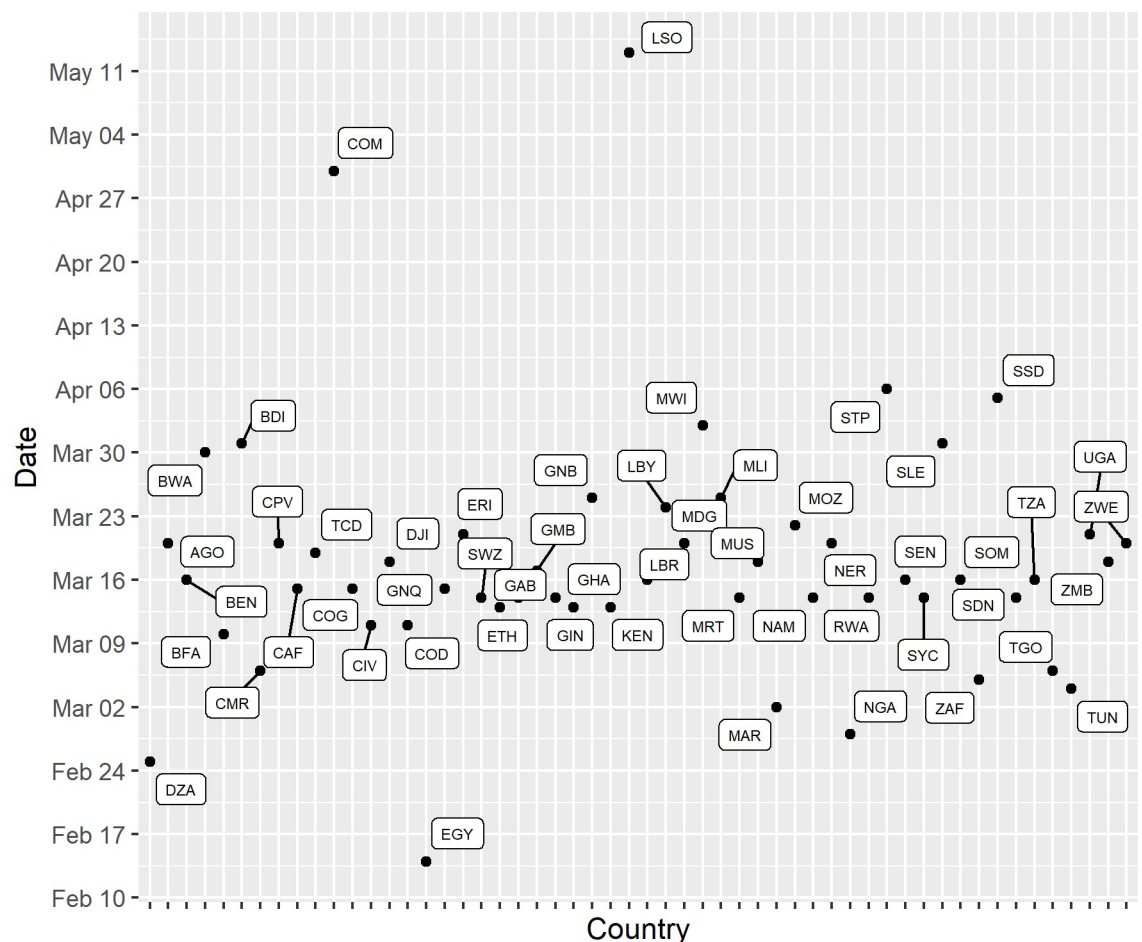


Fig 1. Distribution of the day of the first COVID-19 reported case in 54 African countries. The plot depicts the dates of the first reports of COVID-19 cases in the 54 studied African countries as reported by the Johns Hopkins global time series on the pandemic (CRC, 2020; Dong et al, 2020). The countries in this distribution are designated by their three-letter Alpha-3 codes: DZA: Algeria, AGO: Angola, BEN: Benin, BWA: Botswana, BFA: Burkina Faso, BDI: Burundi, CPV: Cabo Verde, CMR: Cameroon, CAF: Central African Republic, TCD: Chad, COM: Comoros, COD: Democratic Republic of the Congo, COG: Congo, CIV: Côte d'Ivoire, DJI: Djibouti, EGY: Egypt, GNQ: Equatorial Guinea, ERI: Eritrea, SWZ: Eswatini, ETH: Ethiopia, GAB: Gabon, GMB: Gambia, GHA: Ghana, GIN: Guinea, GNB: Guinea-Bissau, KEN: Kenya, LSO: Lesotho, LBR: Liberia, LBY: Libya, MDG: Madagascar, MWI: Malawi, MLI: Mali, MRT: Mauritania, MUS: Mauritius, MAR: Morocco, MOZ: Mozambique, NAM: Namibia, NER: Niger, NGA: Nigeria, RWA: Rwanda, STP: Sao Tome and Principe, SEN: Senegal, SYC: Seychelles, SLE: Sierra Leone, SOM: Somalia, ZAF: South Africa, SSD: South Sudan, SDN: Sudan, TZA: United Republic of Tanzania, TGO: Togo, TUN: Tunisia, UGA: Uganda, ZMB: Zambia, ZWE: Zimbabwe.

<https://doi.org/10.1371/journal.pone.0269573.g001>

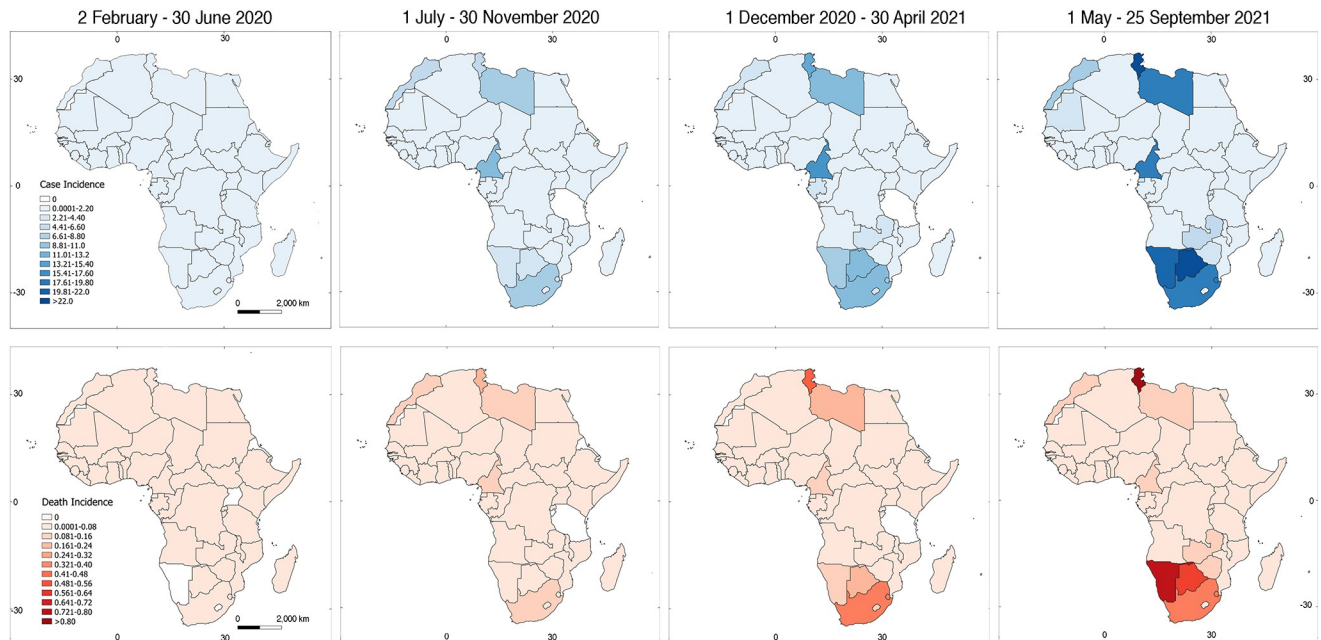


Fig 2. Average case and death incidences of COVID-19 per 100,000 people over four five-month time periods in Africa. Eight plots show average case incidences (upper panels) and average death incidences (bottom panels) over four five-month time periods from 2 February 2020 to 25 September 2021. Scale is the same for all case/death incidence maps and is depicted in the left panels; numbers are individuals affected per 100,000 people.

<https://doi.org/10.1371/journal.pone.0269573.g002>

period = 37.8), Namibia (third period = 8.55; fourth period = 20.3), and South Africa (third period = 9.28; fourth period = 15.7) ranking top among the countries studied (Fig 2). Tanzania had an incidence of 0 for the second and third time periods, which will be discussed below.

COVID-19 death incidence was recorded for all the African countries in the second time period except for Eritrea, Seychelles, Comoros, Mauritius, Tanzania, and Burundi, although the latter four reported 5.51×10^{-3} , 4.83×10^{-3} , 2.39×10^{-4} , and 5.62×10^{-5} death incidences per 100,000 people during the first period, respectively. Further, South Africa (0.219 deaths per 100,000 people) and Tunisia (0.179) reported the highest death incidence in the second period. For the third period, highest death incidences were again reported in South Africa (0.385) and Tunisia (0.422); for the fourth period, highest incidences were recorded in Tunisia (0.808), Namibia (0.732), and Seychelles (0.584).

Few countries lacked information for one or two of the chosen English terms (6/54; 11.1%); only 'coronavirus' and 'COVID19' always recovered search query counts. Several countries that had French or Portuguese listed as an official language returned no information for either one or both language-specific terms (8/32; 25%, S2 Table). Overall, the adjusted R^2 values collected to depict the relationship between GHT search queries and COVID-19 weekly incidence were low, never above 0.4 for any of the countries in either the basic regression or the weighted regression (Fig 3). The largest adjusted R^2 results from the basic regression were for Algeria (0.33), Ethiopia (0.20), and Kenya (0.19; Fig 4). The countries with the lowest adjusted R^2 results from the basic regression included Burkina Faso (-0.028), Sierra Leone (-0.030), and Sudan (-0.031; Figs 3 and 4, S2 Table). For the weighted regression analysis on the first-order differenced case incidence and GHT data, the countries that returned the largest adjusted R^2 results were Guinea-Bissau (0.24), Lesotho (0.08), and Niger (0.07, Fig 3), respectively. The lowest adjusted R^2 results came from Zimbabwe, Egypt, and Mauritania each with an adjusted R^2 value of -0.05 (rounded; see S1 Fig).

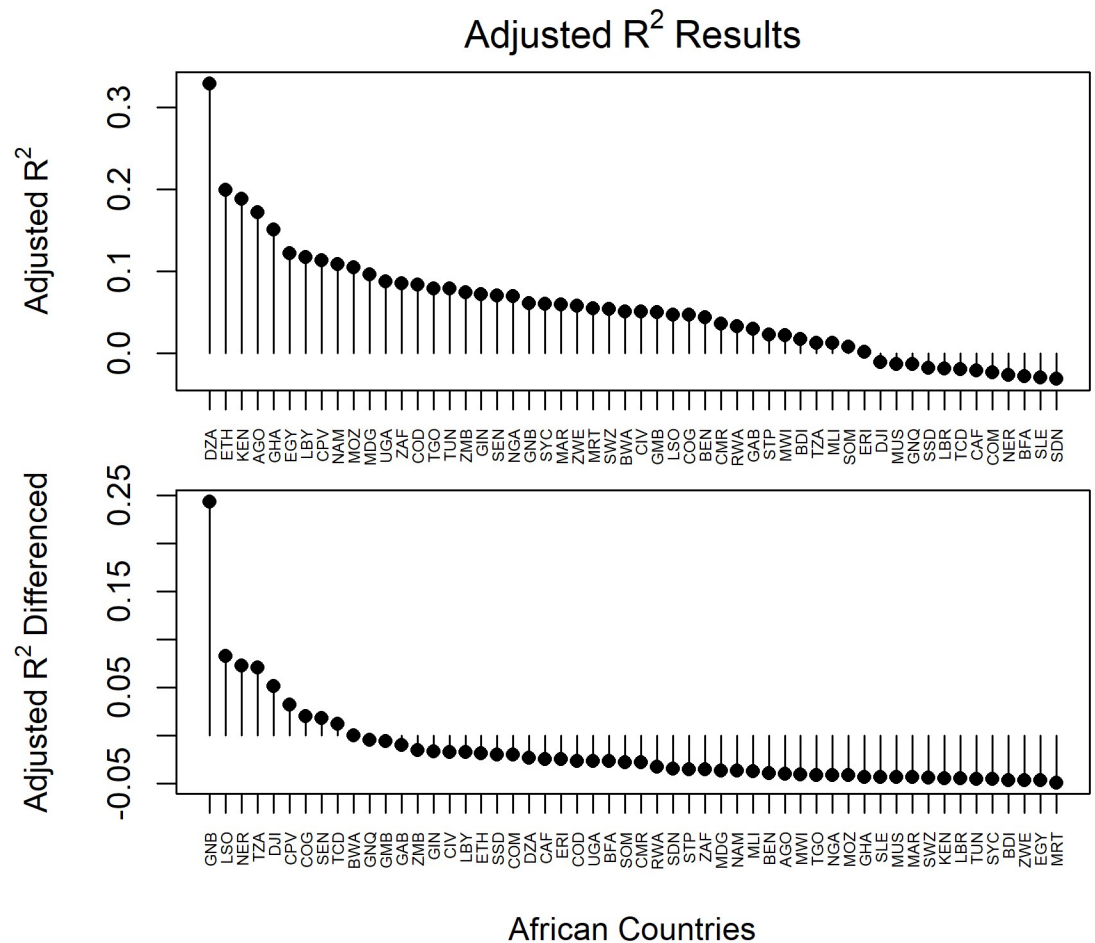


Fig 3. Results of multiple linear regression analysis between COVID-19 incidence and Google Health Trends (GHT) search terms. The adjusted R^2 of the basic (upper panel) and the weighted (bottom panel) regression analysis is depicted here to visually represent the countries from the highest to lowest performance. The countries in this figure are designated by their three-letter Alpha-3 codes as in Fig 1.

<https://doi.org/10.1371/journal.pone.0269573.g003>

Several of the 72 variables were correlated, at least in part, with the pattern of adjusted R^2 statistics obtained for the 54 African countries. Almost all univariate, linear analyses from the basic regression yielded adjusted R^2 values of 0.25 or less, except for the log of average weekly cases (0.37), log of cumulative total deaths (0.30), and log of fixed total number of broadband subscriptions in a country (0.26, S3 Table). The only adjusted R^2 value greater than 0.25 from the weighted regression analyses came from the number of community health workers per 1,000 people, albeit that variable was only available for 26 countries. The log of average weekly cases, the log of GDP, and the log of the volatility scores yielded low adjusted R^2 values ($R^2 = 0.120, 0.057, \text{ and } 0.080$, respectively). The stepwise regression analysis on the untransformed data showed that a model including percentage of GDP for current health expenditure, life expectancy (years) at birth, mobile cellular subscriptions per 100 people, total population, GDP per capita, percentage of people using the Internet, total urban population, total number of mobile cellular subscriptions in a country, average weekly cases over the studied period, and (notably) volatility score for a country calculated using weekly incidence yielded an adjusted

R^2 of 0.40. LASSO regression analysis revealed that a model with life expectancy (years) at birth showed an adjusted R^2 of 0.13.

Examining the adjusted R^2 values collected from the weighted regression model, the volatility score for a country calculated using weekly incidence was deemed most useful for both stepwise and LASSO regressions, though the latter method also returned percentage of the population with access to electricity as an important variable ($R^2 = 0.051$, and 0.063 , respectively). Conversely, using logarithmically transformed variables, a stepwise regression model including average weekly cases over the studied period, percentage of GDP for current health expenditure, life expectancy (years) at birth, yielded an adjusted R^2 value of 0.47. Using the adjusted R^2 values collected from the weighted regression analyses, a stepwise regression model selected the average weekly cases over the time period studied as an important variable. LASSO regression analysis of logarithmic transformed variables indicated that a model including percentage of individuals with access to electricity, life expectancy (years), average weekly cases over the period studied, cumulative deaths, percentage of GDP for current healthcare expenditure, and total population gave the highest adjusted R^2 of 0.45. Using the adjusted R^2 collected from the weighted regression analyses, volatility score for a country calculated using weekly incidence and average weekly cases were returned an $R^2 = 0.13$. The results of these models yielded adjusted R^2 values larger than most of the univariate analyses (S3 Table); thus, although we are mentioning the variables depicting the largest association with GHT and COVID-19 incidence in Africa outputs, we are cautious in interpreting any of the variables as explanatory considering lower than $R^2 = 0.5$ across all models [25,42,43].

Discussion

Despite successful demonstrations of the GHT algorithm to aid infectious disease surveillance for influenza, dengue, and other diseases [26,43,44], our study demonstrates that, in the context of the COVID-19 epidemic, GHT appeared to be difficult to implement as a surveillance tool for COVID-19 incidence and impact. Average weekly cases over the period studied was an important variable when analyzing possible patterns in the adjusted R^2 values collected from both the basic regression and weighted regression analyses. The volatility score for a country was also an important variable for the applicability of GHT, as demonstrated in our univariate, stepwise, and LASSO models (S3 Table). Finally, indicators related to Internet access (mobile cellular subscriptions per 100 people, total number of mobile cellular subscriptions in a country, percentage of individuals using the Internet, percentage of individuals with access to electricity), health (life expectancy (years) at birth), demographics (total population, total urban population), and economics (percentage of GDP for current health expenditure, GDP, GDP per capita) can be interpreted as important factors in the patterns of GHT and COVID-19 incidence although heterogeneously with different modeling approaches and therefore difficult to interpret (S3 Table).

The top three ranking countries based on adjusted R^2 values in the basic regression (Algeria, Ethiopia, and Kenya) all seemed to have similar COVID-19 incidence signal type (Fig 4, upper panels). Cases begin at zero, spike, and subsequently drop to a lower, but still higher level of incidence, followed by additional waves, potentially reflecting an exhaustion of susceptible individuals or dynamics of new variants [45,46]. Algeria, Ethiopia, and Kenya all had strong responses to initial outbreaks of COVID-19 and invested significantly in preventative measures against COVID-19 such as testing, vaccination, and healthcare [47–49]. These three countries also ranked within the top 10 when looking at the total number of mobile cellular subscriptions in a country and GDP (S1 Table). Conversely, Burkina Faso, Sierra Leone, and Sudan are lower-income countries, and have struggled to combat COVID-19 [50–52];

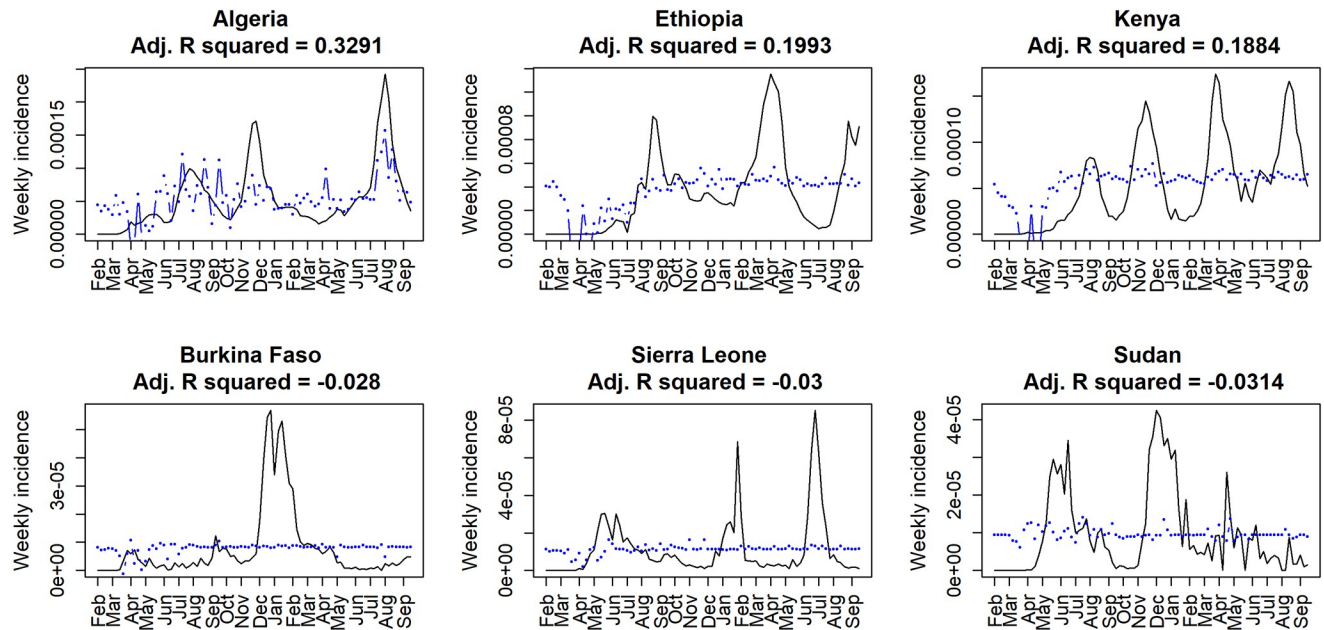


Fig 4. Best and worst performing countries from the basic regression analysis of Google Health Trends (GHT) data. When analyzing whether GHT correlated with case incidence (black line) via a multiple linear regression analysis (blue line), the three best performing countries were Algeria (DZA), Ethiopia (ETH), and Kenya (KEN), respectively (upper panels). The three worst performing countries were Burkina Faso (BFA), Sierra Leone (SLE), and Sudan (SDN), respectively (bottom panels).

<https://doi.org/10.1371/journal.pone.0269573.g004>

according to World Bank data, they ranked lower than the top-ranking countries in terms of total number of mobile cellular subscriptions in a country (S1 Table). Furthermore, while four out of these six countries had an extremely low percentage of individuals using the Internet (< 20% as of 2017), we found no clear association between GHT patterns and Internet accessibility variables, which may indicate that the way Internet access is currently measured reflects GHT behavior poorly. Interestingly, the three countries with the lowest GHT R^2 values (Fig 4, lower panels) showed fewer cases and greater variability in their incidence signal compared to the best-performing countries. The combination of these results may indicate that some consistent level of infection is required for keeping the interest of communities searching information through Google search engines. This gives GHT a chance to match cases, and it may perform better when a rapid growth of infection coincides with interest in the topic and Internet search volume for disease-specific terms is likely to be high, regardless of the level of Internet access.

As in the rest of the world, incidence of both COVID-19 cases and COVID-19 related deaths increased across Africa steadily during the study period. However, in the second and third periods of our study, Tanzania showed zero COVID-19 cases (Fig 2). Upon closer examination, the country stopped reporting coronavirus cases and deaths in April of 2020, so any patterns that might be observed for Tanzania (0.012) are actually reflecting a lack of data [53]. While COVID-19 numbers are concerning on the continent, Africa has been observed to have a lower disease burden in comparison to other regions of the world [54–57, but see 58]. As of August 2020, Africa had reported approximately 69.2 cases and 1.31 deaths per 100,000 people in nearly seven months since COVID-19 was declared a pandemic. For comparison, the U.S. at that point in time, had seen roughly 1500 cases per 100,000 people, and Brazil confirmed roughly 47.0 deaths per 100,000 people [2,3].

Although infodemiology approaches represent the next frontier of infectious disease surveillance [19,23], the present modeling effort demonstrates that search queries from GHT are difficult to correlate with incidence of disease in the context of an emerging epidemic. In contrast with diseases such as influenza or dengue that are studied consistently in a seasonal pattern or are endemic to multiple regions [25,43,59], COVID-19 represented an unprecedented case study that might render Google-based information mining ineffective for several reasons: (a) partial or incomplete COVID-19 case detection and reporting [8,60], (b) media-induced search behavior [61], or even (c) information fatigue [36]. Thus, we encourage caution regarding interpretation of COVID-19 modeling experiments based on Google search engines. For example, Ahmad et al. (2020) found a correlation between gastrointestinal search terms obtained through GT and COVID-19 cases and suggested that Internet searches may be useful in predicting COVID-19 cases using a four-week lag in the U.S. [32]. This correlation, however, might be an artifact since none of the gastrointestinal terms is specific to COVID-19, and the only COVID-19 specific term—‘ageusia’—increased during the time that the pandemic was declared (i.e., 11 March) and decreased while cases started to increase (Fig 1 in [32]). The U.S. showed an increase in case numbers driven by increasing test capacity, thus, these case numbers were reflecting disease incidence inaccurately [62]. Thus, although our findings are based on the GHT algorithm, we are cautious about interpreting our results and those of others in characterizing COVID-19 via Google search engines. Similar to our findings, Asseo et al. (2020) found correlations between GT search queries related to smell and taste at the beginning of the pandemic in Italy and the U.S., which faded in succeeding epidemiological weeks [36]. More importantly, Asseo et al. (2020) also showed how correlation patterns break down when analyzing Google search queries and COVID-19 incidence in nonconsecutive weeks (e.g. 11–17 March vs. 1–7 April 2020 in [36]).

We acknowledge some limitations of the present research. Because of the timeframe of the study and the availability of GHT data as weekly counts, we had to convert daily cases to weekly cases, limiting our analysis to only 86 observations, decreasing the statistical power of our approach. Moreover, the four terms related with COVID-19 that were selected might not be as popular in the region as expected. Language might be an important although permeable barrier [25,27]. Still, in the present study the addition of French and Portuguese translations of search terms did not yield significantly higher adjusted R^2 values (S2 Table). We did not explore the role of media coverage of COVID-19 in web search behavior in Africa, which might be an important cofounder for infodemiology studies [63], however the lack of GHT and COVID-19 associations found in the present research actually demonstrate that even in the context of well-covered epidemics, GHT should be used with caution. Finally, we lacked complete data for some of the variables explored (e.g., prevalence of severe food insecurity in the population; S3 Table) which halts interpretation of several of the indicators used; however, those that were available for all the countries showed certain explanatory power as in other research studies (e.g., total population, signal volatility, disease incidence, etc) [25,43].

Conclusions

Surveillance for an ongoing epidemic via GHT might be useful in specific situations in which accurate case counts can be retrieved and there is sustained level of disease incidence as in the case of dengue or influenza; surveillance via GHT for COVID-19 in Africa seems difficult to implement. Google instruments to recover population search counts—GT and GHT—are potentially powerful digital epidemiology tools that can lead to greater insight into disease dynamics, and should be studied and implemented depending on the particular context of an outbreak [25,30,64–67]. Future directions to examine GHT on COVID-19 research include

expansion of the analysis to a larger dataset both in time and space. Other refinements can be implemented, for example combining other forms of digital data (e.g., Twitter, Wikipedia) to determine if addition of more information improves the predictive power of the model.

Supporting information

S1 Fig. Plots depicting the best and worst performing countries in the weighted regression analysis and plots depicting the multiple linear regression models between COVID-19 case counts and Google Health Trends search queries for the 54 African countries studied in the present manuscript.

(DOCX)

S1 Table. Raw data for the variables used to explore patterns of Google Health Trends search queries and COVID-19 incidence in 54 African countries.

(XLSX)

S2 Table. Date of the first COVID-19 case reported in Africa. Results of multiple linear regression analysis performed between COVID-19 incidence and Google Health Trends search queries from four selected terms.

(XLSX)

S3 Table. Univariate and multivariate linear regression analysis to explore associations between adjusted R^2 of Google Health Trends search queries and COVID-19 incidence in 54 African countries.

(XLSX)

Acknowledgments

DRA thanks the writing-original-group (WOG) for support during the development of this paper.

Author Contributions

Conceptualization: Daniel Romero-Alvarez.

Data curation: Alexander Fulk, Daniel Romero-Alvarez, Qays Abu-Saymeh.

Formal analysis: Alexander Fulk, Daniel Romero-Alvarez.

Funding acquisition: Daniel Romero-Alvarez, A. Townsend Peterson.

Investigation: Alexander Fulk, Daniel Romero-Alvarez, A. Townsend Peterson, Folashade B. Agosto.

Methodology: Alexander Fulk, Daniel Romero-Alvarez, A. Townsend Peterson.

Project administration: Daniel Romero-Alvarez, A. Townsend Peterson, Folashade B. Agosto.

Supervision: Folashade B. Agosto.

Validation: Jarron M. Saint Onge.

Writing – original draft: Alexander Fulk, Daniel Romero-Alvarez, Qays Abu-Saymeh, Jarron M. Saint Onge, A. Townsend Peterson, Folashade B. Agosto.

Writing – review & editing: Alexander Fulk, Daniel Romero-Alvarez, Qays Abu-Saymeh, Jarron M. Saint Onge, A. Townsend Peterson, Folashade B. Agosto.

References

1. CDC. Coronavirus Disease 2019 (COVID-19). 2020a. <https://www.cdc.gov/coronavirus/2019-ncov/need-extra-precautions/older-adults.html> (accessed 3 November 2020).
2. Johns Hopkins Coronavirus Resource Center (CRC). COVID-19 Dashboard. 2020. <https://coronavirus.jhu.edu/> (accessed 7 December 2021).
3. Dong E, Du H, Gardner L. An interactive web-based dashboard to track COVID-19 in real time. *Lancet Infect Dis* 2020; 20:533–34. [https://doi.org/10.1016/S1473-3099\(20\)30120-1](https://doi.org/10.1016/S1473-3099(20)30120-1) PMID: 32087114
4. Williamson EJ, Walker AJ, Bhaskaran K, Bacon S, Bates C, Morton CE, et al. Factors associated with COVID-19-related death using OpenSAFELY. *Nature* 2020; 584:430–6. <https://doi.org/10.1038/s41586-020-2521-4> PMID: 32640463
5. CDC. Coronavirus disease (COVID-19): how is it transmitted? 2020b. <https://www.who.int/news-room/q-a-detail/coronavirus-disease-covid-19-how-is-it-transmitted> (accessed 7 December 2020).
6. Li R, Pei S, Chen B, Song Y, Zhang T, Yang W, et al. Substantial undocumented infection facilitates the rapid dissemination of novel coronavirus (SARS-CoV-2). *Science* 2020; 368:489–93. <https://doi.org/10.1126/science.abb3221> PMID: 32179701
7. Zhang R, Li Y, Zhang AL, Wang Y, Molina MJ. Identifying airborne transmission as the dominant route for the spread of COVID-19. *Proc Natl Acad Sci* 2020; 117:14857–63. <https://doi.org/10.1073/pnas.2009637117> PMID: 32527856
8. Loembé MM, Tshangela A, Salyer SJ, Varma JK, Ouma AEO, Nkengasong JN. COVID-19 in Africa: the spread and response. *Nat Med* 2020; 26:999–1003. <https://doi.org/10.1038/s41591-020-0961-x> PMID: 32528154
9. Campbell F, Archer B, Laurenson-Schafer H, Jinnai Y, Konings F, Batra N, et al. Increased transmissibility and global spread of SARS-CoV-2 variants of concern as at June 2021. *Euro Surveill*. 2021. pii:2100509. <https://doi.org/10.2807/1560-7917.ES.2021.26.24.2100509> PMID: 34142653
10. Mayo Clinic. Herd Immunity and COVID-19 (coronavirus): What you need to know. 2021. <https://www.mayoclinic.org/diseases-conditions/coronavirus/in-depth/herd-immunity-and-coronavirus/art-20486808> (accessed 2 February 2022).
11. Beaubien J. For the 36 countries with the lowest vaccination rates, supply isn't the only issue. *NPR* 2022. <https://www.npr.org/sections/goatsandsoda/2022/01/14/1072188527/for-the-36-countries-with-the-lowest-vaccination-rates-supply-isnt-the-only-issu> (accessed 2 February 2022).
12. Sidibé M. Vaccine Inequity: Ensuring Africa is not left out. *Brookings* 2022. <https://www.brookings.edu/blog/africa-in-focus/2022/01/24/vaccine-inequity-ensuring-africa-is-not-left-out/> (accessed 29 January 2022).
13. Adepoju P. As COVID-19 vaccines arrive in Africa, Omicron is reducing supply and increasing demand. *Nature Medicine* 2021. <https://doi.org/10.1038/d41591-021-00073-x> PMID: 34903893
14. Africa CDC. COVID-19 Vaccination. 2022. <https://africacdc.org/covid-19-vaccination/> (accessed 29 January 2022).
15. Md SI, Tonmoy S, Sazzad HK, Abu-Hena MK, S. M MH, Alamgir K, et al. COVID-19-Related infodemic and its impact on public health: a global social media analysis. *Am J Trop Med Hyg* 2020; 103:1621–29. <https://doi.org/10.4269/ajtmh.20-0812> PMID: 32783794
16. WHO. Key lessons from Africa's COVID-19 vaccine rollout. 2021. <https://www.afro.who.int/news/key-lessons-africas-covid-19-vaccine-rollout> (accessed 29 January 2022).
17. Fairchild G, Tasseff B, Khalsa H, Generous N, Daughton AR, Velappan N, et al. Epidemiological data challenges: planning for a more robust future through data standards. *Front Public Health* 2018; 6:336. <https://doi.org/10.3389/fpubh.2018.00336> PMID: 30533407
18. Wu SL, Mertens AN, Crider YS, Nguyen A, Pokpongkiat NN, Djajadi S, et al. Substantial underestimation of SARS-CoV-2 infection in the United States. *Nat Commun* 2020; 11:4507. <https://doi.org/10.1038/s41467-020-18272-4> PMID: 32908126
19. Salathé M. Digital epidemiology: what is it, and where is it going? *Life Sci Soc Policy* 2018; 14:1–5. <https://doi.org/10.1186/s40504-017-0065-7> PMID: 29302758
20. Vlainac H. Epidemiology, aims and scope. In: Kirch W, editor. *Encyclopedia of public health*. Dordrecht: Springer Netherlands; 2008, p. 350–2.
21. Copeland P, Romano R, Zhang T, Hecht G, Zigmund D, Stefansen C. Google disease trends: an update. *Int. Soc. Neglected Trop. Dis.* 2013. Available from: <https://storage.googleapis.com/pub-tools-public-publication-data/pdf/41763.pdf> (accessed 23 March 2021).
22. Nsoesie EO, Buckeridge DL, Brownstein JS. Guess who's not coming to dinner? Evaluating online restaurant reservations for disease surveillance. *J Med Internet Res* 2014; 16:e22. <https://doi.org/10.2196/jmir.2998> PMID: 24451921

23. Pollett S, Althouse BM, Forshey B, Rutherford GW, Jarman RG. Internet-based biosurveillance methods for vector-borne diseases: Are they novel public health tools or just novelties? *PLoS Negl Trop Dis* 2017a; 11:e0005871. <https://doi.org/10.1371/journal.pntd.0005871> PMID: 29190281
24. Lazer D, Kennedy R, King G, Vespignani A. The parable of Google Flu: traps in big data analysis. *Science* 2014; 343:1203–5. <https://doi.org/10.1126/science.1248506> PMID: 24626916
25. Romero-Alvarez D, Parikh N, Osthus D, Martinez K, Generous N, del Valle S, et al. Google Health Trends performance reflecting dengue incidence for the Brazilian states. *BMC Infect Dis* 2020; 20:252. <https://doi.org/10.1186/s12879-020-04957-0> PMID: 32228508
26. Arora VS, McKee M, Stuckler D. Google Trends: opportunities and limitations in health and health policy research. *Health Policy* 2019; 123:338–41. <https://doi.org/10.1016/j.healthpol.2019.01.001> PMID: 30660346
27. Althouse BM, Ng YY, Cummings DAT. Prediction of dengue incidence using search query surveillance. *PLoS Negl Trop Dis* 2011; 5:e1258. <https://doi.org/10.1371/journal.pntd.0001258> PMID: 21829744
28. Xu C, Bai Y, Yang H, Cao X, Jia P, Wang Y. Detecting lung cancer trends by leveraging real-world and Internet-based data: infodemiology study. *J Med Internet Res*. 2020; 22:e1684. <https://doi.org/10.2196/16184> PMID: 32163035
29. Zhang Q, Perra N, Perrotta D, Tizzoni M, Paolotti D, Vespignani A. Forecasting seasonal influenza fusing digital indicators and a mechanistic disease model. *Proc. 26th Int. Conf. World Wide Web* 2017; 311–19.
30. Kurian SJ, Bhatti A ur R, Alvi MA, Ting HH, Storlie C, Wilson PM, et al. Correlations between COVID-19 cases and Google Trends data in the United States: a state-by-state analysis. *Mayo Clin Proc* 2020; 95:2370–81. <https://doi.org/10.1016/j.mayocp.2020.08.022> PMID: 33164756
31. Brodeur A, Clark A, Fleche S, Powdthavee N. COVID-19, lockdowns and well-being: evidence from Google Trends. *J Public Econ* 2021; 193:104346. <https://doi.org/10.1016/j.jpubeco.2020.104346> PMID: 33281237
32. Ahmad I, Flanagan R, Staller K. Increased Internet search interest for GI symptoms may predict COVID-19 cases in US hotspots. *Clin Gastroenterol Hepatol* 2020; 18:2833–34. <https://doi.org/10.1016/j.cgh.2020.06.058> PMID: 32629121
33. Center for Systems Science and Engineering (CSSE). COVID-19 data repository. Johns Hopkins University. 2020. <https://github.com/CSSEGISandData/COVID-19> (accessed 22 January 2022).
34. Vandenbroucke J, Pearce N. Incidence rates in dynamic populations. *Int J Epidemiol* 2012; 41:1472–79. <https://doi.org/10.1093/ije/dys142> PMID: 23045207
35. United States Census Bureau. International Data Base (IDB). 2020. https://www.census.gov/data-tools/demo/idb/-/country?YR_ANIM=2020&FIPS_SINGLE=AG (accessed 15 December 2020).
36. Asseo K, Fierro F, Slavutsky Y, Frasnelli J, Niv MY. Tracking COVID-19 using taste and smell loss Google searches is not a reliable strategy. *Sci Rep* 2020; 10:20527. <https://doi.org/10.1038/s41598-020-77316-3> PMID: 33239650
37. Nations Online. Official and spoken languages of African countries. https://www.nationsonline.org/oneworld/african_languages.htm (accessed 22 January 2022).
38. Mundfrom D, Smith M, Kay L. The effect of multicollinearity on prediction in regression models. *GLMJ* 2018; 44:24–28.
39. The World Bank. Indicators. 2020. <https://data.worldbank.org/indicator> (accessed 7 December 2020).
40. R Core Team. R: A language and environment for statistical computing. R Foundation for Statistical Computing, Vienna, Austria. 2017. <https://www.R-project.org/> (accessed 7 December 2020).
41. WHO. COVID-19 Public Health Emergency of International Concern (PHEIC) Global research and innovation forum. 2020. [https://www.who.int/publications/m/item/covid-19-public-health-emergency-of-international-concern-\(pheic\)-global-research-and-innovation-forum](https://www.who.int/publications/m/item/covid-19-public-health-emergency-of-international-concern-(pheic)-global-research-and-innovation-forum) (accessed 5 January 2021).
42. James G, Witten D, Hastie T, Tibshirani R. An introduction to statistical learning with applications in R. 1st ed. New York: Springer; 2013.
43. Gluskin R, Johansson M, Santillana M, Brownstein J. Evaluation of Internet-based dengue query data: Google Dengue Trends 2014. *PLoS Negl Trop Dis* 2011; 8:e2713.
44. Messina JP, Brady OJ, Golding N, Kraemer MUG, Wint GRW, Ray SE, et al. The current and future global distribution and population at risk of dengue. *Nat Microbiol* 2019; 4:1508–15. <https://doi.org/10.1038/s41564-019-0476-8> PMID: 31182801
45. Kastalskiy IA, Pankratova EV, Mirkes EM, Kazantsev VB, Gorban AN. Social stress drives the multi-wave dynamics of COVID-19 outbreaks. *Sci Reports* 2021; 11:22497. <https://doi.org/10.1038/s41598-021-01317-z> PMID: 34795311

46. Kupferschmidt K. Viral mutations may cause another 'very, very bad' COVID-19 wave, scientists warn. *Science Insider* 2021. <https://www.science.org/content/article/viral-mutations-may-cause-another-very-very-bad-covid-19-wave-scientists-warn> (accessed 29 January 2022).
47. Klouche-Djedid SN, Shah J, Khodor M, Kacimi SEO, Islam SM, Aiash H. Algeria's response to COVID-19: an ongoing journey. *Lancet* 2021; 9:449.
48. Tesfay F. How conflict has made COVID-19 a neglected epidemic in Ethiopia. *The Conversation* 2021. <https://theconversation.com/how-conflict-has-made-covid-19-a-neglected-epidemic-in-ethiopia-167499> (accessed 29 January 2022).
49. Wangari EN, Gichuki P, Abuor AA, Wambui J, Okeyo SO, Oyatsi HTN, et al. Kenya's response to the COVID-19 pandemic: a balance between minimising morbidity and adverse economic impact. *AAS Open Res.* 2021; 4:3. <https://doi.org/10.12688/aasopenres.13156.2> PMID: 33709055
50. Traoré IT, Ouedraogo S, Kania D, Kaboré FN, Konaté B, Médah R, et al. COVID-19 epidemiological, sociological and anthropological investigation: study protocol for a multidisciplinary mixed methods research in Burkina Faso. *BMC Infect Dis* 2021; 21:896. <https://doi.org/10.1186/s12879-021-06543-4> PMID: 34479501
51. Parmley LE, Hartsough K, Eleeza O, Bertin A, Sesay B, et al. COVID-19 preparedness at health facilities and community service points serving people living with HIV in Sierra Leone. *PLoS ONE* 2021; 16(4):e0250236. <https://doi.org/10.1371/journal.pone.0250236> PMID: 33857253
52. Mukhtar MM, Khogali M. The accelerating COVID-19 epidemic in Sudan. *Nat Immunol* 2021; 22:797–98. <https://doi.org/10.1038/s41590-021-00950-0> PMID: 34035525
53. Makoni M. Tanzania refuses COVID-19 vaccines. *Lancet* 2021; 397:566. [https://doi.org/10.1016/S0140-6736\(21\)00362-7](https://doi.org/10.1016/S0140-6736(21)00362-7) PMID: 33581810
54. Soy A. Coronavirus in Africa: five reasons why Covid-19 has been less deadly than elsewhere. *BBC News* 2020a. <https://www.bbc.com/news/world-africa-54418613> (accessed 22 March 2021).
55. Mwai P. Coronavirus: What's happening to the numbers in Africa? *BBC News* 2020b. https://www.bbc.com/news/world-africa-53181555?intlink_from_url=https://www.bbc.com/news/world&link_location=live-reporting-story (accessed 7 December 2020).
56. European Centers Disease Prevention and Control. COVID-19 situation update worldwide, as of 14 December 2020. 2020. <https://www.ecdc.europa.eu/en/geographical-distribution-2019-ncov-cases> (accessed 15 December 2020).
57. Maeda JM, Nkengasong JN. The puzzle of the COVID-19 pandemic in Africa. *Science* 2021; 371:27–8. <https://doi.org/10.1126/science.abf8832> PMID: 33384364
58. Mwananyanda L, Gill CJ, MacLeod W, Kwenda G, Pieciak R, Mupila Z, et al. Covid-19 deaths in Africa: prospective systematic postmortem surveillance study. *BMJ* 2021; 372:n334. <https://doi.org/10.1136/bmj.n334> PMID: 33597166
59. Pollett S, Boscardin WJ, Azziz-Baumgartner E, Tinoco YO, Soto G, Romero C, et al. Evaluating Google Flu Trends in Latin America: important lessons for the next phase of digital disease detection. *Clin Infect Dis* 2017b; 64:34–41.
60. Schneider EC. Failing the test—the tragic data gap undermining the US pandemic response. *N Engl J Med* 2020; 383:299–302. <https://doi.org/10.1056/NEJMp2014836> PMID: 32412704
61. Southwell BG, Dolina S, Jimenez-Magdaleno K, Squiers LB, Kelly BJ. Zika virus—related news coverage and online behavior, United States, Guatemala, and Brazil. *Emerg Infect Dis* 2016; 22:1320–1. <https://doi.org/10.3201/eid2207.160415> PMID: 27100826
62. Stokes EK, Zambrano LD, Anderson KN, Marder EP, Raz KM, Felix SEB, et al. Coronavirus disease 2019 case surveillance—United States, January 22–May 30, 2020. *Morb Mortal Wkly Rep* 2020; 69:759–65. <https://doi.org/10.15585/mmwr.mm6924e2> PMID: 32555134
63. Sato K, Mano T, Iwata A, Toda T. Need of care in interpreting Google Trends-based COVID-19 infodemiological study results: potential risk of false-positivity. *BMC Med Res Methodol* 2021; 21:1–10.
64. Nuti SV, Wayda B, Ranasinghe I, Wang S, Dreyer RP, Chen SI, et al. The use of Google Trends in health care research: a systematic review. *PLoS ONE* 2014; 9:e109583. <https://doi.org/10.1371/journal.pone.0109583> PMID: 25337815
65. Pelat C, Turbelin C, Bar-Hen A, Flahault A, Valleron A-J. More diseases tracked by using Google Trends. *Emerg Infect Dis* 2009; 15:1327–8. <https://doi.org/10.3201/eid1508.090299> PMID: 19751610
66. Samaras L, García-Barriocanal E, Sicilia MA. Comparing social media and Google to detect and predict severe epidemics. *Sci Rep* 2020; 10:4747. <https://doi.org/10.1038/s41598-020-61686-9> PMID: 32179780
67. Valdivia A, Monge-Corella S. Diseases tracked by using Google Trends, Spain. *Emerg Infect Dis* 2010; 16:168. <https://doi.org/10.3201/eid1601.091308> PMID: 20031078